

Tail risk in government bond markets and ECB unconventional policies*

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São Paulo
August 2017

*Not necessarily the views of ECB or Riksbank.

Contributions

We develop a novel observation-driven model to estimate the time-varying tail shape of time series data from an arbitrary fat-tailed distribution.

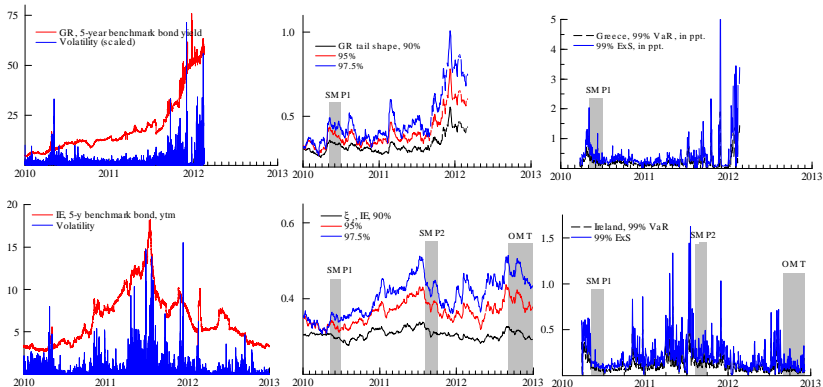
Tail shape dynamics are driven by the score of the predictive log-likelihood. Estimation is straightforward.

The model works reliably in a variety of simulation settings, and for different approaches to the modeling of non-tail observations.

ECB unconventional monetary policies, specifically the Securities Markets Programme (SMP), lowered the inventory risk of dealers making sovereign bond markets during the euro area debt crisis between 2010-2012.

Example: Tail shape and market risk estimates

For Greek & Irish 5-year benchmark bonds, 15 min frequency...



Modeling time-varying tail shape

We introduce time-variation into the tail shape parameter $\xi_t > 0$ of the Generalized Pareto Distribution (GPD).

The pdf of a GPD random variable $x_t = y_t - \tau > 0$ is

$$p(x_t; \delta, \xi_t) = \frac{1}{\delta} \left(1 + \xi_t \frac{x_t}{\delta}\right)^{-\frac{1}{\xi_t} - 1},$$

where $t = 1, \dots, T$, and x_t is the so-called peak-over-threshold (POT) of y_t over a predetermined value τ .

If $y_t - \tau \leq 0$, we consider x_t as missing, or censored (to come).

Setting/assumptions

1. We assume the pdf of y_t , $g(y_t)$, is fat-tailed with time-varying tail-index $\alpha_t > 0$. For example, a Student's t distribution with $\nu_t = \alpha_t = 1/\xi_t$.

Then the CDF $G(y_t)$ can be expressed approximately as

$$G(y_t) = G(\tau) + (1 - G(\tau)) P(x_t)$$

for high values of τ . Tail behavior is captured by the Generalized Pareto distribution $P(x_t; \delta, \xi_t)$.

2. Assume pre-filtered data; fixes $\delta = 1$. Alternative one can treat a bivariate case of t.v. volatility and tail index, for instance Massacci (2015).

Score-driven tail shape

Following Creal, Koopman, and Lucas (2013), Harvey (2013), we endow ξ_t with so-called score-driven dynamics.

To ensure $\xi_t > 0$, we model $f_t = \ln(\xi_t)$. The transition dynamics for f_t are

$$f_{t+1} = \omega + \sum_{i=0}^{p-1} a_i s_{t-i} + \sum_{j=0}^{q-1} b_j f_{t-j},$$

where $s_t = S_t \nabla_t$ is the scaled score

$$\nabla_t = \partial \ln p(x_t; \delta, \xi_t) / \partial f_t$$

$$S_t = E_{t-1} [\nabla_t^2 | f_t, x_{t-1}, x_{t-2}, \dots]^{-1} \text{ scaling function.}$$

Optimality: The score-driven updates always reduce the local Kullback-Leibler divergence between the true conditional density and the model implied conditional density, irrespective of the severity of model misspecification. See Blasques, Koopman, and Lucas (2015).

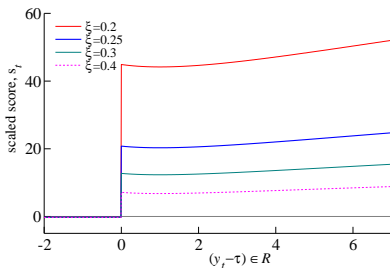
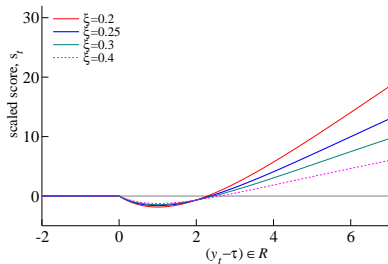
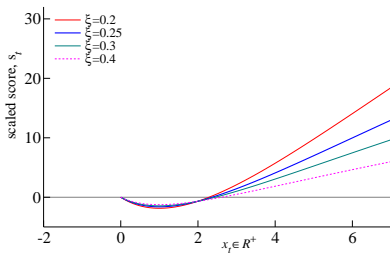
Treatment of non-tail observations

We consider three approaches to handling observations for which $y_t - \tau \leq 0$.

- (1) Deletion of missing entries in x_t , with $t = 1, \dots, T$.
- (2) Model as missing without information about the tail.
- (3) Our favorite approach: model as a draw from a mixture distribution of GPD and a point mass at zero.

Each approach yields different expressions for the score ∇_t and the scaling function S_t .

News impact curves



► Equations

Simulation experiments: Main lessons

- Overall, our score-driven GPD model reliably captures tail shape variation in a variety of simulation settings.
- The simple deletion of missings is appropriate if a complete time series of tail shape estimates is not required, and the tail is sufficiently fat.
- Modeling non-tail observations as missing works well iff there is strong mean reversion in tail shape.
- Modeling via the mixture distribution is most efficient. It works well if the tail is less fat, and also if mean reversion is less strong.
- (Slightly) mis-specified models still work well.

How do asset purchases impact yields?

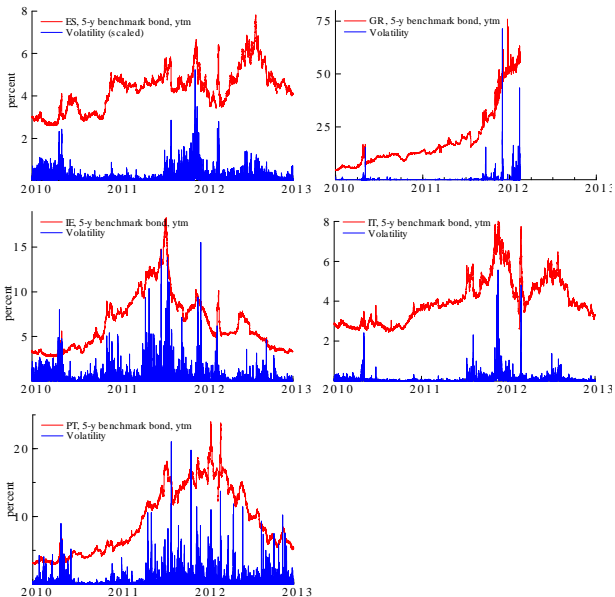
To which extent did ECB non-standard policies impact the tail shape and market risk of euro area sovereign bonds?

Elevated tail risks alone can force institutional investors and market makers to retreat; see Vayanos and Vila (2009), Adrian and Shin (2010).

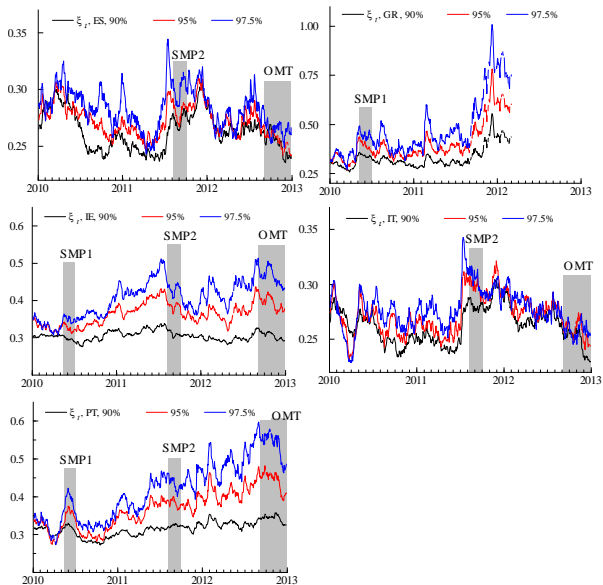
Pelizzon, Subrahmanyam, Tomio, and Uno (2013) provide anecdotal evidence that market makers withdrew from trading Italian debt securities in 2011.

Data: High-frequency bond yields between 2010-2012 for ES, GR, IE, ES, PT. Five-year benchmark bonds from Thomson/Reuters. Sampled at 15 minute frequency as in Ghysels, Idier, Manganelli, and Vergote (2016).
SMP purchase data as in Ghysels et al. (2016).

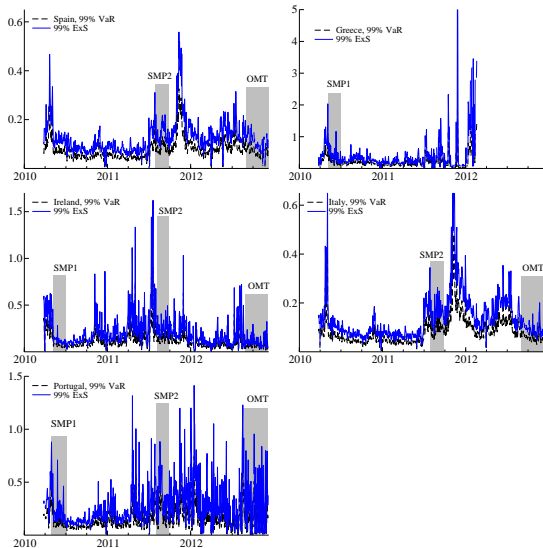
Benchmark bond yields & volatility...



... with corresponding tail shape estimates ...



... and ExS, VaR as market risk measures



Impact identification

The score-driven dynamics for the tail shape parameter can be extended to include contemporaneous or lagged economic variables as additional conditioning variables.

$$f_{t+1} = \omega + a \cdot S_t \nabla_t + b \cdot f_t + c' X_t,$$

where c is a vector of coefficients, and additional variables are in X_t .

We study SMP interventions in government debt securities markets in five euro area countries. Approximately €214 billion (bn) of bonds were acquired between 2010 and early 2012. The SMP announcement dates were 10 May 2010 and 8 August 2011.

Tail shape impact

Tail(0.975) model with GAS- t filtering							
	ω	a	b	SMP_t	$D_{1,t}$	$D_{2,t}$	LogLik
ES	-0.001 (0.000) [0.000]	0.002 (0.000) [0.000]	0.999 (0.000) [0.000]	0.013 (0.006) [0.022]	-0.006 (0.003) [0.036]	-0.006 (0.003) [0.035]	-11672.1
GR	-0.002 (0.001) [0.034]	0.006 (0.001) [0.000]	0.999 (0.001) [0.000]	-0.023 (0.013) [0.068]	0.006 (0.005) [0.259]	0.000 (0.004) [0.954]	-8361.9
IE	-0.001 (0.000) [0.098]	0.004 (0.001) [0.000]	1.000 (0.000) [0.000]	-0.031 (0.035) [0.380]	-0.005 (0.005) [0.312]	0.000 (0.003) [0.962]	-11428.1
IT	-0.001 (0.001) [0.573]	0.002 (0.001) [0.024]	1.000 (0.001) [0.000]	0.004 (0.003) [0.227]	-0.009 (0.003) [0.002]	-0.009 (0.004) [0.036]	-11815.2
PT	-0.002 (0.001) [0.148]	0.003 (0.001) [0.001]	0.999 (0.001) [0.000]	-0.006 (0.025) [0.818]	-0.005 (0.005) [0.339]	-0.001 (0.003) [0.855]	-12124.2

Note: Parameter estimates for the dynamic tail model under the mixture distribution approach. ES, GR, IE, IT, and PT refer to Spanish, Greek, Irish, Italian, and Portuguese five-year bond yields (in 15mins). $D_{1,t}$ and $D_{2,t}$ are announcement dummy variables on 10/05/2010 and 08/08/2011. Standard errors in round brackets and P-values in square brackets.

Conclusion

We developed an easy-to-use statistical model to track the time-varying tail shape and market risk associated with fat-tailed time series data.

ECB unconventional monetary policies lowered the tail shape and market risk of euro area sovereign bonds during the euro area debt crisis between 2010-2012, thereby contributing towards restoring "depth and liquidity" in impaired markets.

Thank you.

Three treatments of non-tail observations.

If $y_t - \tau > 0$ denoted as $I^{(+)}$ doesn't hold, we consider different treatments of x_t :

1. Deletion of missing entries,
2. Missing without information on the tail,
3. A mixture distribution of GPD and a point mass at zero, the density as mentioned before.

Then we can derive the corresponding score-driven dynamics for each treatment.

	Tail index dynamics $\xi_t = \exp(f_t)$	Score and scaling function
1.	$f_{t+1} = \omega + a \cdot S_t \nabla_t + b \cdot f_t$	$\nabla_t = \frac{1}{\xi_t} \ln \left(1 + \xi_t \frac{x_t}{\delta} \right) - (\xi_t + 1) \frac{x_t}{\delta + \xi_t x_t}$ $S_t = (1 + \xi_t)(1 + 2\xi_t)/2\xi_t^2$
2.	$f_{t+1} = \omega + a \cdot I^{(+)} S_t \nabla_t + b \cdot f_t$	$\nabla_t = \frac{1}{\xi_t} \ln \left(1 + \xi_t \frac{x_t}{\delta} \right) - (\xi_t + 1) \frac{x_t}{\delta + \xi_t x_t}$ $S_t = (1 + \xi_t)(1 + 2\xi_t)/2\xi_t^2$
3.	$f_{t+1} = \omega + a \cdot S_t^\phi \nabla_t^\phi + b \cdot f_t$	$\nabla_t^\phi = I^{(+)} \left[\frac{\ln(1+\tau)}{\xi_t} + \nabla_t \right] - I^{(0)} \frac{(1+\tau)^{-\frac{1}{\xi_t}}}{1 - (1+\tau)^{-\frac{1}{\xi_t}}} \frac{\ln(1+\tau)}{\xi_t}$ $S_t^\phi = \left[\frac{2\xi_t^2(1+\tau)^{-\frac{1}{\xi_t}}}{(1+\xi_t)(1+2\xi_t)} + \frac{(1+\tau)^{-\frac{1}{\xi_t}}}{1 - (1+\tau)^{-1/\xi_t}} \frac{\ln(1+\tau)^2}{\xi_t^2} \right]^{-1}$